1 INTRODUCTION

As the core of the modern economy, finance plays an important role in promoting economic growth. On one hand, financial development can give impetus to economic growth directly. On the other hand, financial development can indirectly stimulate economic growth by consumption and investment. As to the research on the relationship between financial development and economic growth, many scholars have their own views. Since the pioneering works of Goldsmith (1969) and Shumpeter (1932), and more recently of McKinnon (1973), Demetriades & Hussein (1996), César Calderón & Lin Liu (2003) and Muhsin Kar, et al. (2011), the relationship between financial development and economic growth has been extensively studied. Demetriades & Hussein (1996) conducted causality tests between financial development and real GDP by using developed time series techniques, and found that there is a causal relationship between financial development and economic growth. César Calderón and Lin Liu (2003) employed the Geweke decomposition test on pooled data to examine the direction of causality between financial development and economic growth, finding that financial development generally leads to economic growth, the Granger causality from financial development to economic growth and the Granger causality from economic growth to financial development coexist. Muhsin Kar, et al. (2011) investigated the direction of causality between financial development and economic growth in the Middle East and North African countries, and the empirical results showed that there is no clear consensus on the direction of causality between financial development and economic growth and, the findings are country specific. All these studies, both theoretical and empirical, have tried to strengthen our understanding of the relationship between financial development and economic growth from different levels.

With the implementation of reform and opening-up, China’s economic growth is very significant, especially since the early 1980s, which is largely due to the reform of its financial system and financial development. Practice development has contributed to the rise of theoretical studies. A great number of theoretical and empirical studies on the relationship between financial development and economic growth in China have emerged. Qi LIANG and Jian-Zhou TENG (2006) explored the relationship between financial development and economic growth based on China’s evidence over the period of 1952 ~ 2001. The empirical results suggested that there exists a unidirectional causality from economic growth to financial development. Chen
(2006) showed that China’s financial development contributes to economic growth, and mobilization of savings and the substitution of loans for budget appropriation are the two channels for the financial sector to contribute to the economy. Guariglia and Poncet (2008) examined the relationship between finance and economic growth by using China’s data from 1989 to 2003, and found that the indicators measuring state intervention in finance are negatively associated with economic growth, but the indicators measuring market-driven financing are positively associated with economic growth. Jin Zhang et al. (2012) investigated the relationship between rural financial development and economic growth at home and abroad, quite a few literatures on the relationship between rural financial development and economic growth are also existed. For example, Zhuang Yan (2011) explored the relationship of rural finance and rural economy in Heilongjiang Province, and the empirical analysis suggested that the low level of rural finance in Heilongjiang province should not have a strong support on the rural economy, but the rural economic development is the Granger causes of rural finance. However, the current literatures on rural financial development and agricultural economic growth are not very systematic, even if there exist a small number of related research literatures. They focus on the role of rural financial development on agricultural economic growth and rarely involve in the agricultural economic impact on rural financial development.

In this paper, we will carry out an empirical study on the relationship between rural financial development and agricultural economic growth based on China’s evidence over the period of 1978-2011.

2 MODEL CONSTRUCTION, METHODOLOGY 
AND DATA DESCRIPTION

2.1 Model construction

When discussing the implementation way of economic growth, the western mainstream economists establish different models of economic growth based on distinct assumptions.

Endogenous economic growth theory is expanded from the classical “Solow model” and includes many specific models, some of which reveal the transmission mechanism of financial development on economic growth. AK model is a simple but very important endogenous economic growth theory. It is based on the AK model that Pagano (1993) has studied the possible ways of financial development affecting economic growth.

In order to clarify the relationship between rural finance and agricultural economy, we introduce a general form of production function, as shown in the following equation:

\[ Y(t) = F(K(t), A(t)L(t)) \]  

Wherein, “Y” denotes time, “Y” represents the total agricultural output, “K” is the capital element, and “AL” is called the effective labor factor. Learning from growth factor analysis approach proposed by Abramovitz (1956) and Solow (1957), we can decompose the equation (1) and obtain the following equation:

\[ Y(t) = \frac{\frac{\partial Y(t)}{\partial K(t)}K(t)}{\frac{\partial Y(t)}{\partial L(t)}L(t)} + A(t)\frac{\partial Y(t)}{\partial A(t)}A(t) \]  

Divide by Y on both sides of equation (2), and then make appropriate adjustments of the above equation, finally we can easily get equation (3):

\[ \frac{Y(t)}{Y(t)} = \frac{\frac{\partial Y(t)}{\partial K(t)}K(t)}{\frac{\partial Y(t)}{\partial L(t)}L(t)} + A(t)\frac{\partial Y(t)}{\partial A(t)}A(t) \]  

In equation (3), \( \alpha_L(t) \) denotes capital-output elasticity at time t; \( \alpha_L(t) \) represents labor-output elasticity of time t; \( R(t) \) is residual value. \( R(t) \) is often referred to as “Solow residual”.

According to the derivation of the above model, we can use the following linear regression model to analyze the impact of rural financial development on agricultural economic growth in China.

\[ y = \beta_0 + \sum_{i=1}^{n} \beta_i x_i + \mu \]  

What needs to be emphasized is that this article investigates the two-way relationship between rural financial development and rural economic growth.

2.2 Methodology and data description

To explore the relationship between rural financial development and rural economic growth in China, we use a Granger-causality test within an ECM model. This test, however, requires that the variables used in a given model be stationary. Many studies have shown that models with non-stationary variables tend to produce spurious regression and make the usual test statistics lack of reliability (Granger & Newbold, 1974). So we will use the ECM model to analyze the rela-
3 EMPIRICAL RESULTS

3.1 Unit root test

Table 1. ADF unit root test results

<table>
<thead>
<tr>
<th>Variables</th>
<th>Test Type (C.T,a)</th>
<th>ADF test statistics</th>
<th>Different significant level 1% 5% 10%</th>
</tr>
</thead>
<tbody>
<tr>
<td>AEG</td>
<td>(C,T,5)</td>
<td>-2.279</td>
<td>-4.324 -3.581 -3.225</td>
</tr>
<tr>
<td>△AEG</td>
<td>(0,0,4)</td>
<td>-6.709</td>
<td>-3.689 -2.972 -2.625</td>
</tr>
<tr>
<td>RFD</td>
<td>(C,T,1)</td>
<td>-2.433</td>
<td>-4.273 -3.558 -3.212</td>
</tr>
<tr>
<td>△RFD</td>
<td>(C,T,0)</td>
<td>-4.906</td>
<td>-4.273 -3.558 -3.212</td>
</tr>
</tbody>
</table>

Note: △AEG and △RFD respectively denote the first-order difference of original series AEG, RFD.

In order to ensure the stationarity of the variables, we adopt standard augmented Dickey-Fuller to make unit root tests on each variable.

The test results (as shown in Table 1) demonstrate that AEG and RFD are non-stationary series. However, △AEG and △RFD are stationary series at all different significant levels. This shows that the original data series are one-order integrated series.

3.2 Cointegration test

Although some time series themselves are non-stationary, the linear combination of them may be steady. According to the unit root test in table one, AEG and RFD are I(1) time series at different significant level. Therefore, we can continue to use the EG two-step method proposed by Engle and Granger (1987) to make cointegration test through soft EViews 6.0.

The first step is to carry out cointegration regression analysis. And then, we can obtain the following regression equation:

\[
AEG = 0.4516 - 0.4397RFD
\]

\[
(30.75) \quad (-10.90)
\]

\[R^2 = 0.7880 \quad F = 7.811 \quad DW = 0.55\]

The EG cointegration theory tells us, if there exists a cointegration relationship between AEG and RFD, then the residual series must be steady. Therefore, we begin to enter the second step analysis.

The second step is to make stationarity test for residual series. The test results on residual series are shown in Table 2:

Table 2. The ADF unit root test of residual series

<table>
<thead>
<tr>
<th>Residual</th>
<th>T-Statistics</th>
<th>P-Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Residual</td>
<td>-3.0289</td>
<td>0.0428</td>
</tr>
</tbody>
</table>

As is known from Table 2, the residual series of regression equation cointegration are stationary under 5% significant level. So there is a long-run integration relationship between AEG and RFD. Equation (7) demonstrates that RFD increases per 1%, AEG will reduce 0.4397% in the long run. Furthermore, an error correction model can be established to analyze the relationship between AEG and RFD.

3.3 Building ECM model

Based on the cointegration relationship between AEG and RFD, we can build the following ECM model:

\[
\Delta AEG_t = -0.006037 - 0.3268\Delta RFD_t - 0.1896 ECM_{t-1}
\]

\[Log = 73.55 \quad AIC = -4.28 \quad SC = -4.14 \]
Seen from the value of Log, AIC and SC, we can conclude that equation (8) has passed the significant test, and the short-run fluctuation of AEG is affected by RFD. The error correction coefficient is negative, which meets the requirement of reverse correction mechanism. The result also shows that there is a negative causal relationship between AEG and RFD in China over the period of 1978 – 2011. In the short-term process of dynamic equilibrium, rural financial development (RFD) changes every 1%, agricultural economic growth (AEG) will change 0.3268% with the opposite direction. The short-term adjustment coefficient is \(-0.1896\), which means that 0.1896% of deviation from the AEG to long-term equilibrium value has been adjusted.

3.4 Granger causality test

The above results show that there is a long-term stable equilibrium relationship between AEG and RFD. But we cannot make sure whether there exists a causal relationship between them or not. Granger causality test can be used to verify the causal relationship between two variables.

Table 3. Granger causality test

<table>
<thead>
<tr>
<th>Null Hypothesis</th>
<th>Lags</th>
<th>F-Statistics</th>
<th>P-Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>RFD does not Granger Cause AEG</td>
<td>2</td>
<td>2.7428</td>
<td>0.0823</td>
</tr>
<tr>
<td>AEG does not Granger Cause RFD</td>
<td>2</td>
<td>1.5083</td>
<td>0.2393</td>
</tr>
</tbody>
</table>

From the test results of Table 3, we can easily find that RFD is a Granger cause of AEG at the 10% significant level, but AEG is not a Granger cause of RFD at the same significant level. It means that rural financial development (RFD) plays an important supporting role in agricultural economic growth (AEG) in China over the period of 1978-2011, but agricultural economic growth (AEG) is not necessarily an important reason for rural financial development (RFD) at the same period.

4 CONCLUSIONS AND POLICY IMPLICATIONS

This paper examines the relationship between rural financial development and agricultural economic growth in China, using time series data over the period of 1978 – 2011. We have empirically evaluated the causality between rural financial development and agricultural economic growth, and found that RFD is the Granger cause of AEG while AEG is not the Granger cause of RFD. Overall, we conclude our findings as follows: (1) Rural financial development has a significant impact on agricultural economic growth in China over the over the period of 1978 – 2011, but the latter has a very weak influence on the former. (2) There is a long-run integration relationship between AEG and RFD, and a negative causal relationship also exists between them. (3) In the short-term process of dynamic equilibrium, RFD’s change will cause AEG’s change in the opposite direction. Obviously, our findings have an important implication of policy recommendation. In China, it is critical to establish a well-developed rural financial system, particularly with sound financial intermediation for rural areas. All of them are very important to the efficient allocation of rural credits, which in turn will help the sustainable growth of agricultural economy.

REFERENCES